

# Benchmarking of a Bayesian single cell RNAseq differential gene expression test for dose-response study designs

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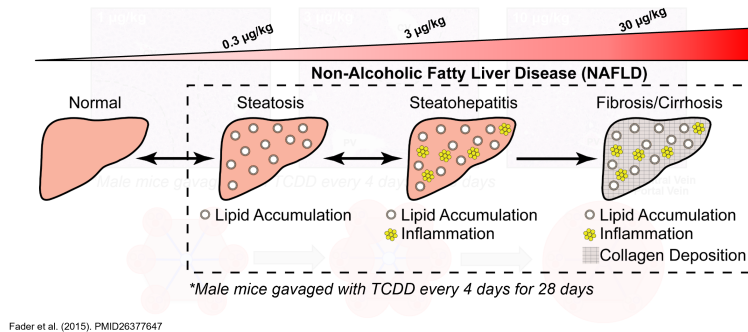
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# Motivation

- 2,3,7,8-Tetrachlorodibenzo-p-Dioxin (TCDD)- a highly toxic and persistent organic pollutant
  - A by-products of a wide range of manufacturing processes including smelting, chlorine bleaching of paper pulp and the manufacturing of some herbicides and pesticides
  - Found in soils, sediments and food, especially dairy products, meat, fish and shellfish
- This contaminant can be found at the **superfund sites**

# Objective

Fader et al. <sup>1</sup> used **bulk RNAseq data** on liver and on the intestine as well as flow cytometry (this is the only cell-specific level data) and histology (involving female mice)



**Goal:** Investigate dose-dependent TCDD-elicited hepatic cell-specific gene expression associated with the development of NAFLD among **male mice** using **scRNAseq data**

<sup>1</sup>Fader et al. (2015). 2,3,7,8-Tetrachlorodibenzo-p-Dioxin Alters Lipid Metabolism and Depletes Immune Cell Populations in the Jejunum of C57BL/6 Mice. Toxicol Sci. 2015 Dec;148(2):567-80.

# Normal versus fatty liver



**Non-Alcoholic Fatty Liver Disease (NAFLD)**



# Experimental Design

- Randomly assigned male C57BL/6 mice to one of the eight levels, 0 with 0.1 mL sesame oil vehicle (0-level or control) 0.01, 0.03, 0.1, 0.3, 1, 3, 10 or 30  $\mu\text{g}/\text{kg}$  TCDD every 4 days for 28 days (How many mice in each dose level? 3, so the total was  $3 \times 8 = 24$ )
- Hepatic single-nuclei RNA-sequencing (snRNAseq) was performed using the 10 $\times$  Genomics Chromium Single Cell 3' v3.1 kit
- These are mRNA measurements

# National Toxicology Program's approach to genomic dose response modeling

- Design dose response (DR) experiment with a sufficient number of doses
- Design appropriate statistical test of hypothesis for deriving genes that exhibit minimum effect to treatment
- Fit parametric DR models derived from the Environmental Protection Agency (EPA) software to identify a biological potency estimate
- Group genes into predefined sets defined by gene ontologies and compute composite **POD** of the gene set
- Provide biological explanations for the selected set of genes and POD estimates

**POD**: Point of departure, the threshold dose level at which the gene expression starts to change from the control group

# Challenges

- Traditional tests for pre-filtering: ANOVA
- scRNA-seq data is highly heterogeneous (**across different cell types**) and has a large number of zeroes
- Violates standard Gaussian assumptions
- No recommended method for differential gene expression analysis (DGEA) in multiple group single cell experimental studies
- Our contribution: **Bayesian multiple group test (scBT)**
  - Designed exclusively for dose–response scRNAseq data
  - FDR control

# Distributional Assumptions

- $Y_{i,j,k}$ : expression value of cell  $i$ , gene  $j$ , dose-level  $k$ , for  $i = 1, \dots, n$  and  $j = 1, \dots, p$ ,  $k = 1, \dots, K$
- $R_{i,j,k} = I[Y_{i,j,k} > 0]$ : indicator denoting the presence of an expression
- Adopt the Hurdle model <sup>2</sup>

$$\begin{aligned} Y_{i,j,k} | R_{i,j,k} = 1 &\sim \text{Normal}(\mu_{j,k}, \sigma_j^2), \\ \text{pr}(Y_{i,j,k} = 0 | R_{i,j,k} = 0) &= 1, \\ R_{i,j,k} &\sim \text{Bernoulli}(\omega_{j,k}), \end{aligned} \tag{1}$$

- $\mu_{j,k}$ : the mean expression of the  $j$ th gene, level  $k$ , when it is expressed
- $\omega_{j,k}$ : the rate of gene expression of gene  $j$  and dose-level  $k$

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<sup>2</sup>McDavid et al. (2013). Data exploration, quality control and testing in single-cell qPCR-based gene expression experiments. *Bioinformatics*, 29, 461–467.



# Test of Hypothesis

- Need to account for the bimodality (due to the mixing of zeros and the positive valued numeric variable) in single cell gene expression distribution
- Design a test capable of detecting changes in the means and the inflation parameters, simultaneously, across the dose levels

$$H_{0,j} : \mu_{j,1} = \cdots = \mu_{j,K} = \mu_j \text{ and } \omega_{j,1} = \cdots = \omega_{j,K} = \omega_j$$

versus the alternative

$$H_{a,j} : H_{0,j} \text{ does not hold}$$

# Bayesian Test- priors

For a given gene, say  $j$ , we use the following priors to calculate the marginal likelihoods

$H_{0,j}$	$H_{a,j}$
$\mu_{j,1} = \dots = \mu_{j,K} = \mu_j$ $\mu_j \sim N(m_0, \tau_\mu \sigma_j^2)$	$\mu_{j,k} \sim \text{Normal}(m_{k,0}, \tau_{k,\mu} \sigma_j^2)$
$\sigma_j^2 \sim IG(a_\sigma, b_\sigma)$	$\sigma_j^2 \sim IG(a_\sigma, b_\sigma)$
$\omega_{j,1} = \dots = \omega_{j,K} = \omega_j$ $\omega_j \sim \text{Beta}(a_\omega, b_\omega)$	$\omega_{k,j} \sim \text{Beta}(a_{k,\omega}, b_{k,\omega})$

Hyperparameters are obtained by maximising the marginal likelihood under the null and the alternative hypothesis

# Bayesian Test of Hypothesis (scBT)

- Bayes factor

$$BF_{01,j} = \frac{\mathcal{L}_{H_0,j}}{\mathcal{L}_{H_a,j}} \times \frac{\pi(H_{a,j})}{\pi(H_{0,j})}$$

- $\pi(H_{a,j})$  and  $\pi(H_{0,j})$ : prior probabilities for alternative and null model  
Are  $\pi(H_{a,j}) = \pi(H_{0,j})$ ? Yes.
- $\mathcal{L}_{H_0,j}$  and  $\mathcal{L}_{H_a,j}$ : marginal likelihood under the null and the alternative hypothesis.

# Marginal Likelihood under $H_{0,j}$

$$\begin{aligned} \mathcal{L}_{H_{0,j}} &= \frac{1}{(2\pi)^{(\sum_{k=1}^K \sum_{i=1}^{n_k} R_{i,j,k})/2}} \times \frac{1}{\sqrt{1 + \tau_\mu \sum_{k=1}^K \sum_{i=1}^{n_k} R_{i,j,k}}} \\ &\times \frac{1}{\Gamma(a_\sigma) b_\sigma^{a_\sigma}} \times \frac{\Gamma(a_\sigma + (\sum_{k=1}^K \sum_{i=1}^{n_k} R_{i,j,k})/2)}{(1/b_\sigma + \mathcal{A}_{tot}/2)^{a_\sigma + (\sum_{k=1}^K \sum_{i=1}^{n_k} R_{i,j,k})/2}} \\ &\times \frac{\text{Beta}(a_\omega + (\sum_{k=1}^K \sum_{i=1}^{n_k} R_{i,j,k}), b_\omega + \sum_{k=1}^K n_k - (\sum_{k=1}^K \sum_{i=1}^{n_k} R_{i,j,k}))}{\text{Beta}(a_\omega, b_\omega)} \end{aligned}$$

where

$$\mathcal{A}_{tot} = \left\{ \sum_{k=1}^K \sum_{i=1}^{n_k} R_{k,i,j} Y_{i,j,k}^2 + \frac{m_0^2}{\tau_\mu} \right\} - \left\{ \sum_{k=1}^K \sum_{i=1}^{n_k} R_{i,j,k} + \frac{1}{\tau_\mu} \right\}^{-1} \left\{ \sum_{k=1}^K \sum_{i=1}^{n_k} R_{i,j,k} Y_{i,j,k} + \frac{m_0}{\tau_\mu} \right\}$$

# Marginal Likelihood under $H_{a,j}$

$$\begin{aligned} \mathcal{L}_{H_{a,j}} &= \frac{1}{(2\pi)^{(\sum_{k=1}^K \sum_{i=1}^{n_k} R_{k,i,j})/2}} \times \frac{1}{\prod_{k=1}^K \sqrt{1 + \tau_{k,\mu} \sum_{i=1}^{n_k} R_{i,j,k}}} \\ &\times \frac{1}{\Gamma(a_\sigma) b_\sigma^{a_\sigma}} \times \frac{\Gamma(a_\sigma + \sum_{k=1}^K \sum_{i=1}^{n_k} R_{i,j,k}/2)}{(1/b_\sigma + \sum_{k=1}^K \mathcal{A}_k/2)^{a_\sigma + \sum_{k=1}^K \sum_{i=1}^{n_k} R_{i,j,k}/2}} \\ &\times \prod_{k=1}^K \frac{\text{Beta}(a_{k,\omega} + \sum_{i=1}^{n_k} R_{i,j,k}, b_{k,\omega} + n_k - \sum_{i=1}^{n_k} R_{i,j,k})}{\text{Beta}(a_{k,\omega}, b_{k,\omega})} \end{aligned}$$

where

$$\mathcal{A}_k = \left\{ \sum_{i=1}^{n_k} R_{i,j,k} Y_{i,j,k}^2 + \frac{m_{k,0}^2}{\tau_{k,\mu}} \right\} - \left\{ \sum_{i=1}^{n_k} R_{i,j,k} + \frac{1}{\tau_{k,\mu}} \right\}^{-1} \left\{ \sum_{i=1}^{n_k} R_{i,j,k} Y_{i,j,k} + \frac{m_{k,0}}{\tau_{k,\mu}} \right\}^2$$

# Multiplicity control

- $D_j$ : Data for the  $j$ th gene
- Calculate the posterior probability of the null hypothesis

$$p(H_{0,j}|D_j) = \left(1 + \frac{1}{BF_{01,j}}\right)^{-1}$$

- For a target FDR  $\alpha$ , we reject  $H_{0,j}$  when  $p(H_{0,j}|D_j) < \zeta$ ,<sup>3</sup>
  - $\zeta$  is the largest value such that  $\frac{C(\zeta)}{J(\zeta)} \leq \alpha$
  - $J(\zeta) = \{j : p(H_{0,j}|D_j) \leq \zeta\}$  and  $C(\zeta) = \sum_{j \in J(\zeta)} p(H_{0,j}|D_j)$
  - $\frac{C(\zeta)}{J(\zeta)}$ : the average posterior probability of null hypothesis of the statistically significant genes

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<sup>3</sup>Newton et al. (2004). Detecting differential gene expression with a semiparametric hierarchical mixture method. *Biostatistics*, 5, 155–176.

# Competing approaches

- Likelihood ratio test (LRT-multiple)
- LRT with a linear model for the mean and the inflation parameters (LRT-linear)
  - $\mu_{j,k} = \beta_{0,j} + \beta_{j,1}d_k$  ( $d_k$ :  $k$ th dose,  $j$ th gene)
  - $\text{logit}(\omega_{j,k}) = \gamma_{0,j} + \gamma_{j,1}d_k$
  - For the  $j$ th gene, test  $H_{0,j} : \beta_{j,1} = 0, \gamma_{j,1} = 0$  versus  $H_{a,j} : \beta_{j,1} \neq 0$  or  $\gamma_{j,1} \neq 0$

# Competing approaches (existing)

- limma-trend<sup>4</sup>: linear regression with dose as the explanatory variable
- MAST<sup>5</sup>: Model-based Analysis of Single-cell Transcriptomics
- Seurat Bimod<sup>6</sup>: (a pairwise test assuming the single cell RNA-seq hurdle model framework)
- WRS: Wilcoxon-Rank Sum test (it is pairwise test)
- ANOVA
- KW: Kruskal-Wallis test (nonparametric extension of one-way ANOVA)

For all these methods, we used the [Benjamini-Hochberg adjusted  \$p\$ -value](#)

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<sup>4</sup>Law et al. (2014) voom: Precision weights unlock linear model analysis tools for RNA-seq read counts. *Genome Biol.*, 15, R29.

<sup>5</sup>Finak et al. (2015). MAST: a flexible statistical framework for assessing transcriptional changes and characterizing heterogeneity in single-cell RNA sequencing data. *Genome Biol.*, 16, 278.

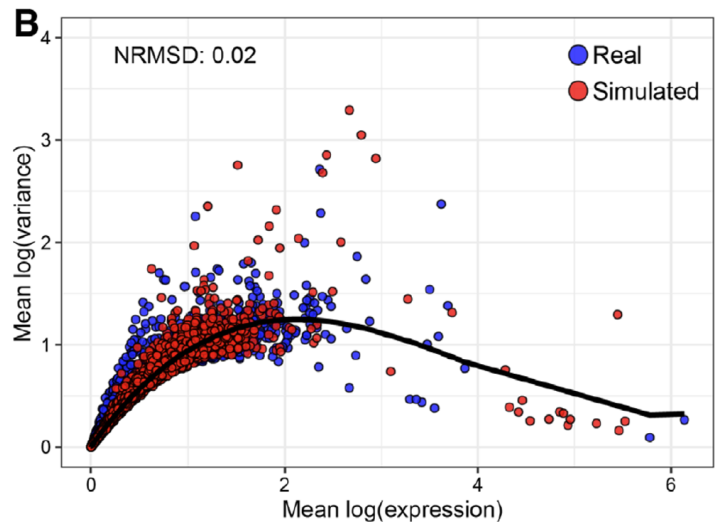
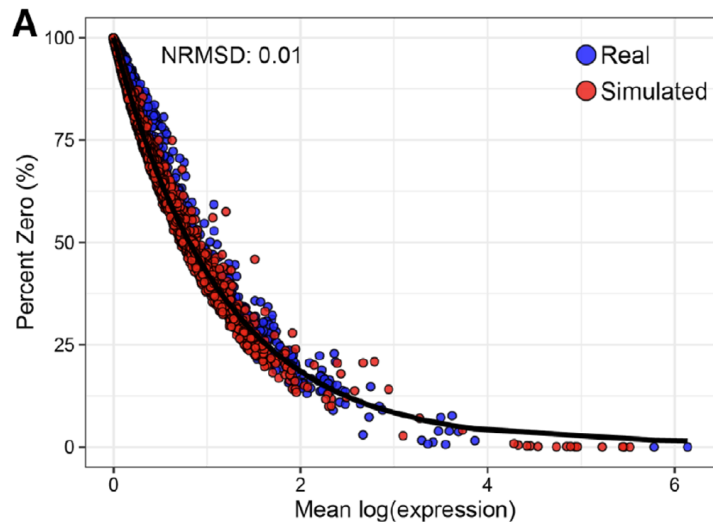
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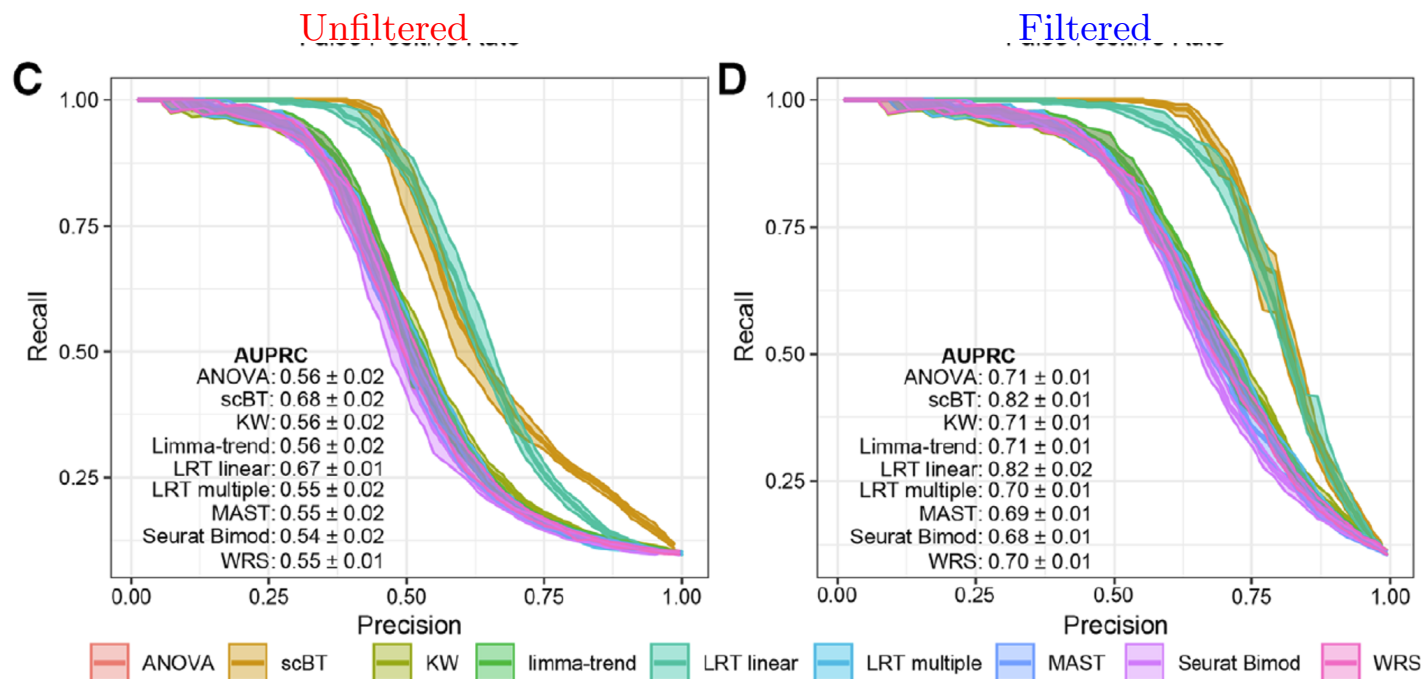
# Simulations

- Data was simulated consisting of nine dose groups of 500 cells each (4500 total) and 5000 genes with a 10% probability of being DE
- Simulation of genes based on initial parameters derived from real DR data was reproduced 10 times
- Performance of scBT was benchmarked against 8 other DE analysis tests
- To investigate test performance in controlling type I errors, DGEA methods on simulated datasets were examined with 0 %DE genes (i.e. negative control)

# Simulations

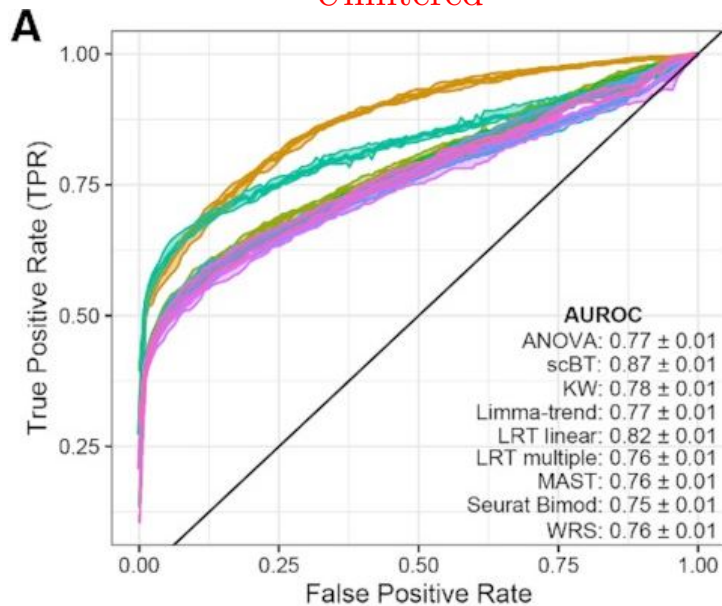


# Simulation Results (PRC)

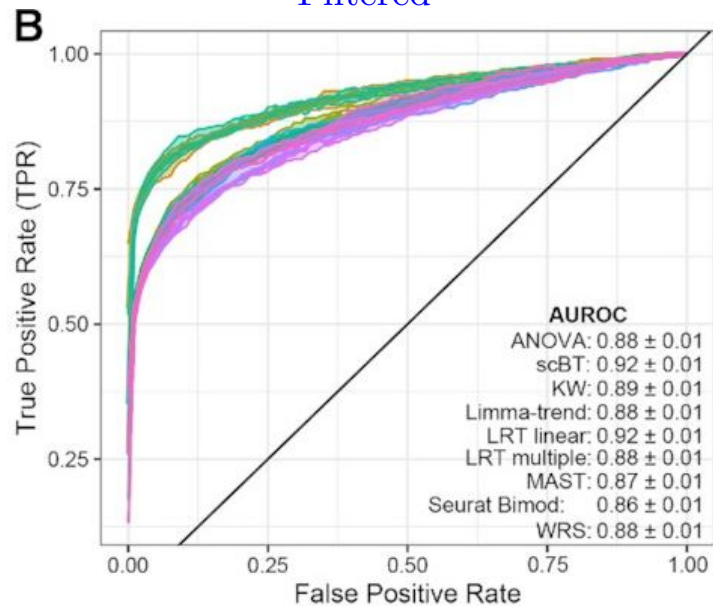


# Simulation Results (ROC)

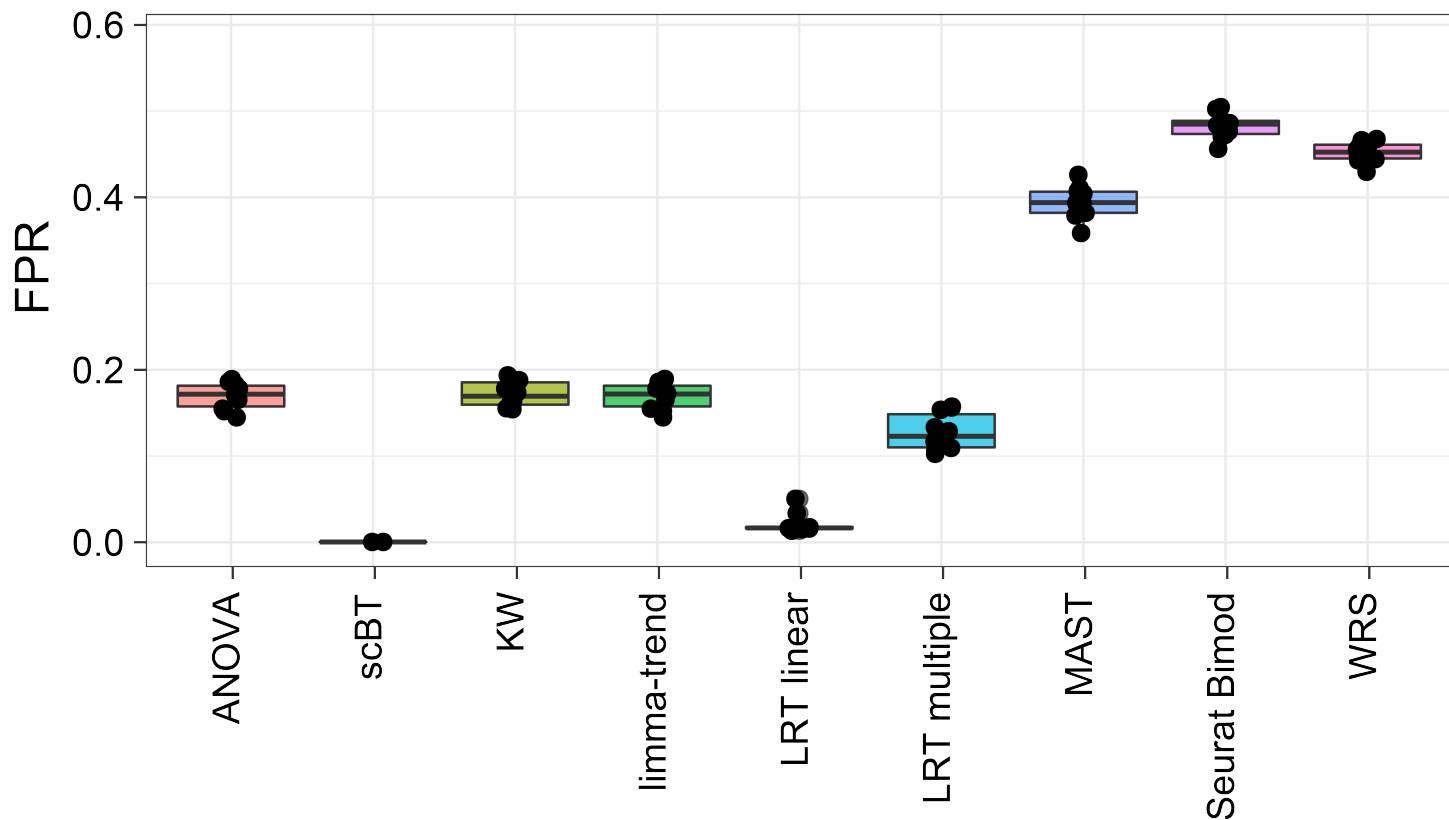
Unfiltered



Filtered



# Simulation Results



# Simulation Results (performance ranking)

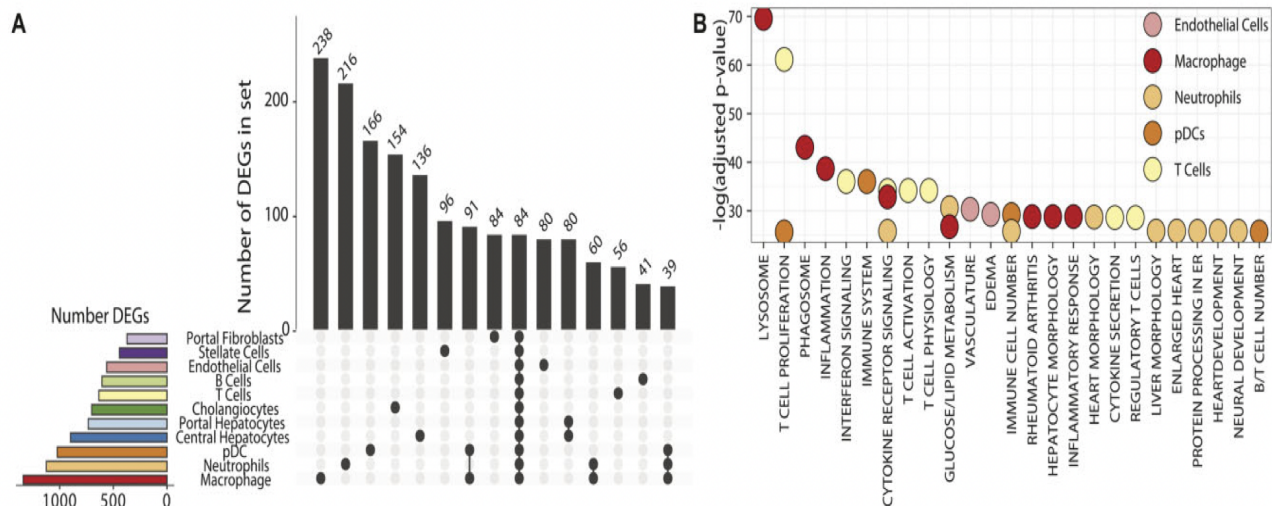
The methods were compared based on the 1) Mathew's correlation coefficients (MCC), 2) false positive rates (FPRs), 3) false negative rates (FNRs), 4) the area under the ROC curve (AUROC) and 5) the area under the precision recall curve (AUPRC)

	Fit-for-purpose		Multiple groups tests				Two group tests		
AUPRC Rank	1	2	5	4	4	6	7	9	7
AUROC Rank	1	2	4	3	5	8	6	9	6
MCC Rank	1	2	5.5	4	5.5	3	7	9	8
FNR Rank	9	7	5.5	4	5.5	7	3	1	2
FPR Rank	1	2	5	6	5	3	7	9	8
Overall Rank	1	2	5	4	5	6	7	9	7
	scBT	LRT linear	ANOVA	KW	limma-trend	LRT multiple	MAST	Seurat Bimod	WRS

# Real Dose Response Dataset

- Total nuclei: 131613
- Average nuclei per dose group: 14624
- Seurat was used to integrate and log-normalize expression data
- Average of 1,665 genes were detected across all nuclei
- Applied the scBT method, used the FDR controlled selection criteria
- **Filtering method:** Genes in the experimental dataset were considered differentially expressed when expressed in  $\geq 5\%$  of cells in at least one dose group and had a  $|\text{fold-change}| \geq 1.5$  in at least one of the seven treatment groups

# Real Data Analysis



**Figure 2.** Set and functional analysis of hepatic differentially expressed genes (DEGs) from male mice gavaged with TCDD every 4 days for 28 days. A, UpSet plot of the 15 largest gene sets, in rank order, based on set analysis that identified both unique and common DEGs (Bayes Factor adjusted FDR  $\geq 0.05$  and  $|\text{fold-change}| \geq 1.5$ ) among all identified cell types. Set sizes represent the number of genes identified in only the cells as indicated by black circles. B, Functional analysis of DEGs for each cell type using gene lists from KEGG and MPO from the Gene Set Knowledgebase (GSKB; <http://ge-lab.org/gskb/>). Gene sets with  $\geq 60\%$  overlap were combined and manually annotated. The top 30 enriched functions (adjusted  $p$ -value) across all cell types are shown. A complete list is available in [Supplementary Table 2](#).



# Quick Summary

- Developed a multiplicity corrected Bayesian multiple group test (scBT), designed exclusively for DGEA of dose–response scRNAseq data
- In the context of investigating chemical or drug MoAs, false positives have the potential to lead to wasted effort and resources
- Simulations: scBT has excellent FPR control and top ranked AUPRC, but scBT has low power
- Real datasets: scBT detected biologically relevant genes in NAFLD development and progression
- The real data are deposited at the Gene Expression Omnibus

- Simulator : **R package SplattDR** is available at <https://github.com/zacharewskil>
- Proposed test approaches : **R package scBT** is available at <https://github.com/satabdisaha1>



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Research article

## Single-cell transcriptomics shows dose-dependent disruption of hepatic zonation by TCDD in mice

Rance Nault <sup>1,2</sup> Satabdi Saha,<sup>3</sup> Sudin Bhattacharya <sup>2,4</sup> Samiran Sinha,<sup>5</sup> Tapabrata Maiti,<sup>3</sup> Tim Zacharewski<sup>1,2,\*</sup>

JOURNAL ARTICLE

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