

This manual illustrates how to use the R package ‘mesub’

This packages estimates the regression parameters $\boldsymbol{\beta} = (\beta_1^T, \beta_2)^T$ of the model

$$H(T) = -\beta_1^T \mathbf{Z} - \beta_2 X + e,$$

when X is not observed in the sample data. Here H is an unknown monotone transformation function, e is a random variable with a known distribution and is independent of \mathbf{Z} and X .

The observed data are n iid copies of $(T^*, \Delta, \mathbf{Z}, \mathbf{W}^*)$, where \mathbf{Z} is a $p \times 1$ vector of observed covariates, $T^* = \min(T, C)$, where T is the time-to-failure, C is the censoring time, and $\Delta = I(T \leq C)$. We assume that conditional on the complete covariates (X, \mathbf{Z}) , where X is unobservable, T and C are independent. Here $\mathbf{W}^* = (W_1^*, \dots, W_m^*)^T$ is a surrogate measurement for the scalar covariate X . We assume that

$$W_{ij}^* = X_i + U_{ij}^*, i = 1, \dots, n, j = 1, \dots, m,$$

where the measurement errors U_{ij}^* are assumed to be iid copies of a random variable U^* which follows a symmetric distribution centered at 0. We assume that U^* is independent of (T, X, \mathbf{Z}, C) .

In the following examples we will simulate data using varying mechanism, and then use the package to get the estimates of the parameters and the standard errors.

```
library(mesub)
set.seed(1)
r=0
m=2
trbeta1=1
trbeta2=1
sigmau=0.71
n=200 #Sample size
ran=runif(n, 0, 1)
nby3=as.integer(0.33*n)
x=c(rnorm(nby3, -0.6, 0.5), rnorm((n-nby3), 1.25, 0.5)) # true covariate x
x=x+6
z=runif(n, 0, 1) # error free covariate
if(r>0) epsilon=log(((1+0)*exp(-r*log(ran)) -1)/r) else epsilon=log(-log(ran))
t=exp(-trbeta1*z-trbeta2*x+epsilon) # this is the actual time to event
cns=runif(n, 0, 6) # 10% censored data with r=0
delta=as.numeric(t<=cns) # right censoring indicator
tstar=apply(cbind(t, cns), 1, min)
ustar=matrix(rnorm((m*n), 0, sigmau), ncol=m, nrow=n) # measurement error
wstar=x+ustar
wstar=wstar/sd(wstar) # It is always better to use rescaled values
z=as.matrix(z)
## End of data generation
out=ltm.me2(tstar, delta, wstar, z, r, 2)
out
$naive.estimate
[1] 1.125509 1.023513
# here 1.125509 is the estimate of  $\beta_1$  and 1.023513
# is the estimate for  $\beta_2$ 
```

```

$naive.vcov
      [,1]      [,2]
[1,] 0.071864708 0.004930186
[2,] 0.004930186 0.009131966
$naive.se
[1] 0.26807594 0.09556132
$our.estimate
[1] 1.444724 1.695173
$our.vcov
      [,1]      [,2]
[1,] 0.11731595 0.02277929
[2,] 0.02277929 0.06662100
$our.se
[1] 0.3425142 0.2581104
$AIC
[1] 151.9277
$BIC
[1] 175.0159
out=ltm.me2(tstar, delta, wstar, z, r, 1)
> out
$naive.estimate
[1] 1.125509 1.023513
$naive.vcov
      [,1]      [,2]
[1,] 0.071864708 0.004930186
[2,] 0.004930186 0.009131966
$naive.se
[1] 0.26807594 0.09556132
$our.estimate
[1] 1.405962 1.708010
$our.vcov
      [,1]      [,2]
[1,] 0.13117580 0.03455378
[2,] 0.03455378 0.07227576
$our.se
[1] 0.3621820 0.2688415
$AIC
[1] 163.8657
$BIC
[1] 173.7606

```

The above results clearly indicate that the estimates due to the two methods are quite different. Furthermore, based on the BIC criteria we should prefer a normal model for X given Z over a mixture model for X given Z. In the next example we simulate measurement errors from an uniform distribution.

```

library(mesub)
set.seed(1)
r=0
m=3
trbeta1=1
trbeta2=1
sigmau=0.71

```

```

n=200 #Sample size
ran=runif(n, 0, 1)
nby3=as.integer(0.33*n)
x=c(rnorm(nby3, -0.6, 0.5), rnorm((n-nby3), 1.25, 0.5)) # true covariate x
z=runif(n, 0, 1) # error free covariate
if(r>0) epsilon=log(( (1+0)*exp(-r*log(ran)) -1)/r) else epsilon=log(-log(ran))
t=exp(-trbeta1*z-trbeta2*x+epsilon) # this is the actual time to event
cns=runif(n, 0, 6) # 10% censored data with r=0
delta=as.numeric(t<=cns) # right censoring indicator
tstar=apply(cbind(t, cns), 1, min)
ustar=matrix(runif((m*n), -1.75, 1.75)*sigmau, ncol=m, nrow=n) # measurement error
wstar=x+ustar
wstar=wstar # It is always better to use rescaled values
z=as.matrix(z)
ptime=proc.time()
out=ltm.me2(tstar, delta, wstar, z, r, 2)
ptime=proc.time()-ptime
ptime
  user system elapsed
23.279  0.001  25.048
out
$naive.estimate
[1] 1.2392133 0.8885428
$naive.vcov
      [,1]      [,2]
[1,] 0.077681536 0.005833456
[2,] 0.005833456 0.008015719
$naive.se
[1] 0.27871408 0.08953055
$our.estimate
[1] 1.391408 1.166867
$our.vcov
      [,1]      [,2]
[1,] 0.11180842 0.01634633
[2,] 0.01634633 0.02605338
$our.se
[1] 0.3343777 0.1614106
$AIC
[1] 229.4864
$BIC
[1] 252.5746

> out=ltm.me2(tstar, delta, wstar, z, r, 1)
out
$naive.estimate
[1] 1.2392133 0.8885428
$naive.vcov
      [,1]      [,2]
[1,] 0.077681536 0.005833456
[2,] 0.005833456 0.008015719
$naive.se
[1] 0.27871408 0.08953055
$our.estimate
[1] 1.424818 1.168084

```

```

$our.vcov
      [,1]      [,2]
[1,] 0.11832461 0.02040846
[2,] 0.02040846 0.02730206
$our.se
[1] 0.3439834 0.1652333
$AIC
[1] 245.47
$BIC
[1] 255.365

```

The following example considers a linear transformation model with $r = 2$ and with four covariates.

```

set.seed(1)
r=2
m=4
trbeta1=0.5
trbeta2=0.5
trbeta3=0.35
trbeta4=0.85
sigmau=1.0
n=500 #Sample size
ran=runif(n, 0, 1)
nby3=as.integer(0.33*n)
x=c(rnorm(nby3, -0.6, 0.5), rnorm((n-nby3), 1.25, 0.5)) # true covariate x
z1=runif(n, 0, 1) # error free covariate
z2=rnorm(n, 0, 1) # error free covariate
z3=rbinom(n, 1, 0.28)# error free covariate
if(r>0) epsilon=log(((1+0)*exp(-r*log(ran)) -1)/r) else epsilon=log(-log(ran))
t=exp(-trbeta1*z1-trbeta2*z2-trbeta3*z3-trbeta4*x+epsilon) # this
# is the actual time to event
cns=runif(n, 0, 6) # 10% censored data with r=0
delta=as.numeric(t<=cns) # right censoring indicator
tstar=apply(cbind(t, cns), 1, min)
ustar=matrix(runif((m*n), -1.75, 1.75)*sigmau, ncol=m, nrow=n) # measurement error
wstar=x+ustar
wstar=wstar # It is always better to use rescaled values
z=cbind(z1, z2, z3)
z=as.matrix(z)
ptime=proc.time()
out=ltm.me2(tstar, delta, wstar, z, r, 2)
ptime= proc.time()-ptime
ptime
  user  system elapsed
262.752   0.013  262.963
  out
$naive.estimate
[1] 0.6437162 0.3406450 0.4619951 0.6426618
# the first three components of naive.estimate are for z1, z2, z3,
# respectively, and 0.6426618 corresponds to x
$naive.vcov
      [,1]      [,2]      [,3]      [,4]
[1,] 0.157511775 2.504083e-03 0.0010257574 2.297990e-03

```

```

[2,] 0.002504083 1.420488e-02 0.0007483502 -6.941956e-06
[3,] 0.001025757 7.483502e-04 0.0671921231 8.035732e-04
[4,] 0.002297990 -6.941956e-06 0.0008035732 1.242336e-02
$naive.se
[1] 0.3968775 0.1191842 0.2592144 0.1114601
$our.estimate
[1] 0.6404138 0.3351174 0.4636812 0.8453742
$our.vcov
      [,1]      [,2]      [,3]      [,4]
[1,] 0.1621392059 2.477169e-03 0.0006852633 2.877967e-03
[2,] 0.0024771690 1.471947e-02 0.0006702920 -5.460864e-06
[3,] 0.0006852633 6.702920e-04 0.0693538997 9.261138e-04
[4,] 0.0028779674 -5.460864e-06 0.0009261138 2.324422e-02
$our.se
[1] 0.4026651 0.1213238 0.2633513 0.1524606
$AIC
[1] 618.5021
$BIC
[1] 664.8628
  out=ltm.me2(tstar, delta, wstar, z, r, 1)
  out
$naive.estimate
[1] 0.6437162 0.3406450 0.4619951 0.6426618
$naive.vcov
      [,1]      [,2]      [,3]      [,4]
[1,] 0.157511775 2.504083e-03 0.0010257574 2.297990e-03
[2,] 0.002504083 1.420488e-02 0.0007483502 -6.941956e-06
[3,] 0.001025757 7.483502e-04 0.0671921231 8.035732e-04
[4,] 0.002297990 -6.941956e-06 0.0008035732 1.242336e-02
$naive.se
[1] 0.3968775 0.1191842 0.2592144 0.1114601
$our.estimate
[1] 0.6627085 0.3373030 0.4706584 0.8262940
$our.vcov
      [,1]      [,2]      [,3]      [,4]
[1,] 0.1620212875 2.427448e-03 0.0007725748 3.328742e-03
[2,] 0.0024274481 1.468441e-02 0.0006527443 4.349151e-05
[3,] 0.0007725748 6.527443e-04 0.0691885159 1.138072e-03
[4,] 0.0033287422 4.349151e-05 0.0011380721 2.215119e-02
$our.se
[1] 0.4025187 0.1211793 0.2630371 0.1488327
$AIC
[1] 613.636
$BIC
[1] 634.709

```